

## 1 Feynman-Kac Formula

Feynman-Kac formula is an important connection between SDE and parabolic PDE. Consider the following boundary value problem on  $[0, T] \times \mathbb{R}$ ,

$$\begin{cases} \frac{\partial F}{\partial t}(t, x) + b(t, x)\frac{\partial F}{\partial x}(t, x) + \frac{1}{2}\sigma^2(t, x)\frac{\partial^2 F}{\partial x^2}(t, x) = 0, \\ F(T, x) = \Phi(x). \end{cases} \quad (1)$$

In general, there is no closed form solution of the above PDE. Feynman-Kac formula says that the solution of the above PDE can be written as the expectation of certain stochastic process, defined by the following SDE defined over  $[t, T]$ :

$$\begin{cases} dX_s = b(s, X_s)ds + \sigma(s, X_s)dW_s, \\ X_t = x. \end{cases} \quad (2)$$

Then one can conclude that

$$F(t, x) = \mathbb{E}_{t,x}[\Phi(X_T)], \quad \text{for } t \leq T$$

which is the Feynman-Kac formula. Here  $\mathbb{E}_{t,x}$  means that the initial point of the stochastic process is  $X_t = x$ .

**Proof:** Consider the infinitesimal generator of the SDE:

$$A = b(t, x)\frac{\partial}{\partial x} + \frac{1}{2}\sigma^2(t, x)\frac{\partial^2}{\partial x^2}.$$

So, Eq. (1) can be rewritten as

$$\begin{cases} \frac{\partial F}{\partial t}(t, x) + AF(t, x) = 0, \\ F(T, x) = \Phi(x). \end{cases} \quad (3)$$

By Ito's Lemma, we have

$$dF = \left(\frac{\partial F}{\partial t} + AF\right)ds + \sigma\frac{\partial F}{\partial x}dW_s = \sigma\frac{\partial F}{\partial x}dW_s. \quad (4)$$

$$F(T, X_T) = F(t, x) + \int_t^T \sigma\frac{\partial F}{\partial x}dW_s. \quad (5)$$

Take the expectation, we have  $F(t, x) = \mathbb{E}_{t,x}[\Phi(X_T)]$ , which is the Feynman-Kac formula.  $\square$

**Komogorov Backward Equation (KBE):** An important corollary of Feynman-Kac formula is that the function  $F(t, x) = P(x_{t'} = x' | x_t = x)$  (for fixed  $t', x'$  and  $t' > t$ ) satisfies the first PDE in Eq equation 1. Namely, the following

$$\frac{\partial P(x_{t'} = x' | x_t = x)}{\partial t}(t, x) + b(t, x) \frac{\partial P(x_{t'} = x' | x_t = x)}{\partial x}(t, x) + \frac{1}{2} \sigma^2(t, x) \frac{\partial^2 P(x_{t'} = x' | x_t = x)}{\partial x^2}(t, x) = 0.$$

This equation is also known as Komogorov Backward Equation (KBE). One way to see this is as follows: for fixed  $t'$  with  $t' = T$ , let  $F(T, x) = \Phi(x) = \delta_{x'}(x)$ . Hence, we can see that

$$F(t, x) = \mathbb{E}_{t,x}[\Phi(X_T)] = \mathbb{E}[\delta_{x'}(X_T) | X_t = x] = P(x_{t'} = x' | x_t = x).$$

## 2 Fokker Planck Equation

There is a closely related equation called Kolmogorov Forward Equation, which is also known as Fokker Planck Equation in physics literature.

### 2.1 Overview

**Definition 1 (Hermitian adjoint)** Each linear operator  $A$  on a Euclidean vector space defines a Hermitian adjoint (or adjoint) operator  $A^*$  on that space according to the rule

$$\langle Ax, y \rangle = \langle x, A^*y \rangle. \quad (6)$$

where  $\langle \cdot, \cdot \rangle$  is the inner product on the vector space.

**Definition 2 (Formal adjoint in one variable)** In the functional space of square-integrable functions on a real interval  $(a, b)$ , the scalar product is defined by

$$\langle f, g \rangle = \int_a^b \overline{f(x)} g(x) dx,$$

where  $\overline{f(x)}$  denotes the complex conjugate of  $f(x)$ . We assume that  $f$  or  $g$  are smooth and their values and derivatives vanish as  $x \rightarrow a, x \rightarrow b$  (such functions are typically called test functions). Consider the differential operator  $T$  ( $T$  maps a function to a function), defined as follows:

$$Tu = \sum_{k=0}^n a_k(x) D^k u.$$

Here  $D$  is the differential operator. One can also define the adjoint of the linear differential operator  $T$  as

$$T^*u = \sum_{k=0}^n (-1)^k D^k [\overline{a_k(x)} u]. \quad (7)$$

It is not difficult to prove that Eq. equation 7 is indeed the adjoint (i.e., it satisfies Definition 1). We only need to apply integral by part ( $\int f dg = fg - \int g df$ ) repeatedly and we leave it as an exercise.

**Theorem 3 (Fokker Planck Equation / Kolmogorov Forward Equation)** Assume that  $b(x)$  is  $C^1$  and  $\sigma(x)$  is  $C^2$ . For  $\rho \in C^2$ , define

$$\mathcal{L}^* \rho(x) = - \sum_{i=1}^n \frac{\partial}{\partial x^i} (b^i(x) \rho(x)) + \frac{1}{2} \sum_{i,j=1}^n \sum_{k=1}^m \frac{\partial^2}{\partial x^i \partial x^j} (\sigma^{ik}(x) \sigma^{jk}(x) \rho(x)). \quad (8)$$

Suppose that the density  $p_t(x)$  exists and is  $C^1$  in  $t$ ,  $C^2$  in  $x$ . Then

$$\frac{\partial}{\partial t} p_t(x) = \mathcal{L}^* p_t(x), t \in [0, T], \quad (9)$$

i.e. the density  $p_t(x)$  of  $X_t$  must satisfy the Fokker Planck Equation (Kolmogorov Forward Equation).

**Proof:** Fix an  $f \in C_0^2$  (in  $C^2$  and with compact support). By Ito's rule, we obtain

$$f(X_t) = f(X_0) + \int_0^t \mathcal{L} f(X_s) ds + \text{martingale}. \quad (10)$$

(the last term is a martingale as  $f$ , and hence its derivatives, have compact support, and thus the integrand is bounded). Taking the expectation and using Fubini's theorem, we obtain

$$\mathbb{E}(f(X_t)) = \mathbb{E}(f(X_0)) + \int_0^t \mathbb{E}(\mathcal{L} f(X_s)) ds. \quad (11)$$

Substituting the definition of  $p_t(y)$ , integrating by parts, and using Fubini's theorem again, we have

$$\int_{\mathbb{R}^n} f(y) p_t(y) dy = \int_{\mathbb{R}^n} f(y) p_0(y) dy + \int_{\mathbb{R}^n} f(y) \int_0^t \mathcal{L}^* p_s(y) ds dy \quad (12)$$

Now note that this expression holds for any  $f \in C_0^2$ , so we can conclude that

$$a(y) = p_t(y) - p_0(y) - \int_0^t \mathcal{L}^* p_s(y) ds = 0 \quad (13)$$

for all  $y$ , except possibly on some subset with measure zero *w.r.t.* the Lebesgue measure.  $\square$

## 2.2 Ornstein-Uhlenbeck Process

**Definition 4 (Ornstein-Uhlenbeck Process)** The Ornstein-Uhlenbeck process  $x_t$  is defined by the following stochastic differential equation:

$$dx_t = -\theta x_t dt + \sigma dW_t \quad (14)$$

where  $\theta > 0$  and  $\sigma > 0$  are parameters and  $W_t$  denotes the Brownian motion process. See some examples in Figure 1. Note that the OU process is a stationary process.

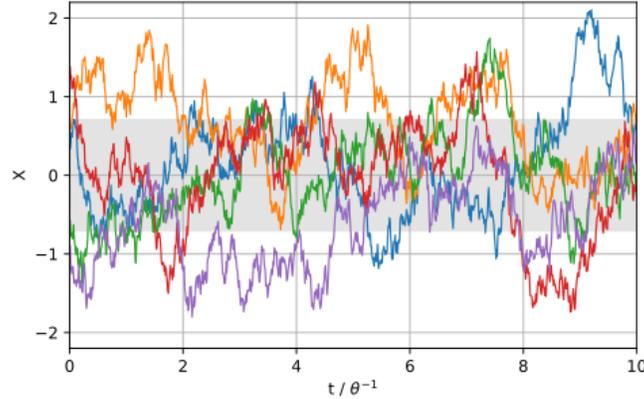


Figure 1: Five simulations with  $\theta = 1, \sigma = 1$

The Ornstein–Uhlenbeck process can also be described in terms of a probability density function,  $P(x, t)$ , which specifies the probability of finding the process in the state  $x$  at time  $t$ . This function satisfies the Fokker–Planck equation

$$\frac{\partial P}{\partial t} = \theta \frac{\partial}{\partial x}(xP) + D \frac{\partial^2 P}{\partial x^2} \quad (15)$$

where  $D = \sigma^2/2$ . This is a linear parabolic partial differential equation which can be solved by a variety of techniques. The transition probability, also known as the Green’s function,  $P(x, t | x', t')$  is a Gaussian with mean  $x' e^{-\theta(t-t')}$  and variance  $\frac{D}{\theta} (1 - e^{-2\theta(t-t')})$ :

$$P(x, t | x', t') = \sqrt{\frac{\theta}{2\pi D(1 - e^{-2\theta(t-t')})}} \exp \left[ -\frac{\theta}{2D} \frac{(x - x' e^{-\theta(t-t')})^2}{1 - e^{-2\theta(t-t')}} \right] \quad (16)$$

This gives the probability of the state  $x$  occurring at time  $t$  given initial state  $x'$  at time  $t' < t$ . Equivalently,  $P(x, t | x', t')$  is the solution of the Fokker–Planck equation with initial condition  $P(x, t') = \delta(x - x')$ .

### 2.3 Heat Equation

Heat equation on Euclidean space is a special case of Fokker-Planck equation.

**Definition 5 (Heat Equation)** *Given an open subset  $U \subseteq \mathbb{R}^n$  and a subinterval  $I \subseteq \mathbb{R}$ , one says that a function  $u : U \times I \rightarrow \mathbb{R}$  is a solution of the heat equation if*

$$\frac{\partial u}{\partial t} = \alpha \nabla^2 u = \alpha \Delta u \quad (17)$$

where  $\Delta$  is the Laplacian operator.

Heat kernel solves the Heat equation. As

$$\frac{\partial K}{\partial t}(x, y, t) = \Delta_x K(x, y, t), \quad (18)$$

we have the solution is

$$K(x, y, t) = \frac{1}{(4\pi t)^{d/2}} e^{-\frac{|x-y|^2}{4t}} \quad (19)$$

with  $\lim_{t \rightarrow 0} K(x, y, t) = \delta(x - y) = \delta_x(y)$ .

## 2.4 Gibbs Distribution

Consider the SDE with the energy field  $E$  (the drift direction is the negative gradient direction):

$$dx_t = -\nabla_x E(x_t) dt + \sqrt{\frac{2}{\beta}} dW_t \quad (20)$$

Let  $p_t(\cdot)$  be the density of  $x_t$ . We want to find the stationary (density) distribution  $p(x)$ . In other words, if the initial distribution  $x_0 \sim p(x)$ ,  $x_t$  is also distributed as  $p(x)$ , i.e.,  $p_t = p$  for all  $t \geq 0$ . By Fokker Planck Equation, we know that

$$\frac{\partial p_t}{\partial t}(x) = \mathcal{L}^* p_t.$$

Since  $p_t$  does not change with  $t$ , so  $\partial p_t / \partial t = 0$ . Equivalently we have  $\mathcal{L}^* p_t = 0$ , or more concretely

$$\nabla^\top (p_t \nabla E) + \frac{1}{\beta} \nabla^\top \nabla p_t = 0. \quad (21)$$

So  $p_t \nabla E + \frac{1}{\beta} \nabla p_t$  should be a constant, say  $C$ . As  $p_t$  should be smooth and integrable, so at infinity  $p_t$  and  $\nabla p_t$  should approach to 0, so the constant  $C = 0$ . Hence, we have

$$\nabla(E + \frac{1}{\beta} \log p_t) = 0.$$

Solving the above equation give the following solution:

$$p_t(x) \propto \exp(-\beta E(x)). \quad (22)$$

A stationary p.d.f.  $p$  of the above form is called the *Gibbs distribution*.

## 3 Reverse Time Diffusion Model

### 3.1 Reverse-time SDE

Much of this section follows Appendix B of (Song et al., 2020), so it might be a good idea to huff it straight from the source now that we have all the tools to understand it. There are a few extra things explicitly derived here, so let us keep moving forward.

We restrict the diffusion coefficient to be a scalar (or a scalar multiplied with the identity matrix) which only depends on the time  $t$  and not  $X$ . The forward SDE is

$$dX_t = f(X_t, t)dt + g(t)dW_t \quad (23)$$

Here  $f$  is a vector function and  $W_t$  is the standard Wiener process with time ranging from 0 to  $T$ . A remarkable theorem of Anderson [1] show that the reverse-time process  $\{X_t^{\leftarrow}\}$  can be also described as an SDE as follows:

$$\begin{aligned} dX_t^{\leftarrow} &= (f(X_t^{\leftarrow}, t) - \frac{g^2(t)}{P_t(X_t^{\leftarrow})} \nabla_x p(X_t^{\leftarrow}))dt + g(t)d\bar{W}_t \\ &= (f(X_t^{\leftarrow}, t) - g^2(t) \nabla_x \log p_t(X_t^{\leftarrow}))dt + g(t)d\bar{W}_t \end{aligned} \quad (24)$$

Here,  $p_t(\cdot)$  is the distribution of  $X_t$  at time  $t$  in the forward process. In the reverse SDE, the time ranges from  $T$  to 0 and thus  $dt$  is a negative increment, and  $d\bar{W}_t$  is the reversed Wiener process (or Gaussian increment with variance  $-dt$ ).

In particular, suppose the initial distribution of  $X_0$  (forward process) is  $p_0$  and the terminal distribution is  $p_T$ . Let the reverse process  $X_t^{\leftarrow}$  starts from the distribution  $p_T$  at time  $T$ . Now, suppose the time goes backwards. Anderson shows that  $X_t^{\leftarrow}$  follows the same distribution as  $X_t$ . A complete proof of this fact can be found in [2].

In fact, more is true: not only the marginal distribution of  $X_t^{\leftarrow}$  (at any time) is the same as that of  $X_t$ , the joint distribution of the sample path  $\{X_t^{\leftarrow}\}$  is the same as that of  $\{X_t\}$ , and it is possible to construct a backward sample path from each forward sample path (they are not the same!). See Anderson's original paper [1] for the details.

## 4 Score-Based Generation Model

We briefly introduce two popular score-based diffusion models, SMLD and DDPM. Both models consist of a forward process and backward process. In the forward process, we start from the data distribution  $X(0)$  and gradually add Gaussian noise until the distribution  $X(T)$  becomes close to pure Gaussian noise. In the backward process, we start from  $X(T)$  and gradually remove the noise and generate  $X(0)$  by simulating the reverse process guaranteed by Anderson's theorem. See Figure 2. Both SMLD and DDPM can be regarded as discretizations of the above forward and reverse SDEs.

### 4.1 Denoising Score Matching With Langevin Dynamics (SMLD) [3]

Let  $p_{data}(x)$  denote the data distribution. Let  $p_\sigma(\tilde{x}|x) := \mathcal{N}(\tilde{x}; x, \sigma^2 I)$  be a perturbation kernel, and  $p_\sigma(\tilde{x}) := \int p_{data}(x)p_\sigma(\tilde{x}|x)dx$ . Consider a sequence of positive noise scales  $\sigma_{\min} = \sigma_1 < \sigma_1 < \dots < \sigma_N = \sigma_{\max}$ . Typically,  $\sigma_{\min}$  is small enough such that  $p_{\sigma_{\min}} \approx p_{data}(x)$ , and  $\sigma_{\max}$  is large enough such that  $p_{\sigma_{\max}} \approx \mathcal{N}(x; 0, \sigma_{\max}^2 I)$ . The perturbation kernel  $p_\sigma(\tilde{x}|x)$  corresponds to the following Markov Chain:

$$x_i = x_{i-1} + \sqrt{\sigma_i^2 - \sigma_{i-1}^2} z_{i-1},$$

where  $z_{i-1} \sim N(0, I)$ . In the limit where  $\sigma_i$  becomes a continuous function  $\sigma(t)$ , the above Markov Chain can be viewed as a discretization of the following SDE:

$$dx = \sqrt{\frac{d[\sigma^2(t)]}{dt}} dw.$$

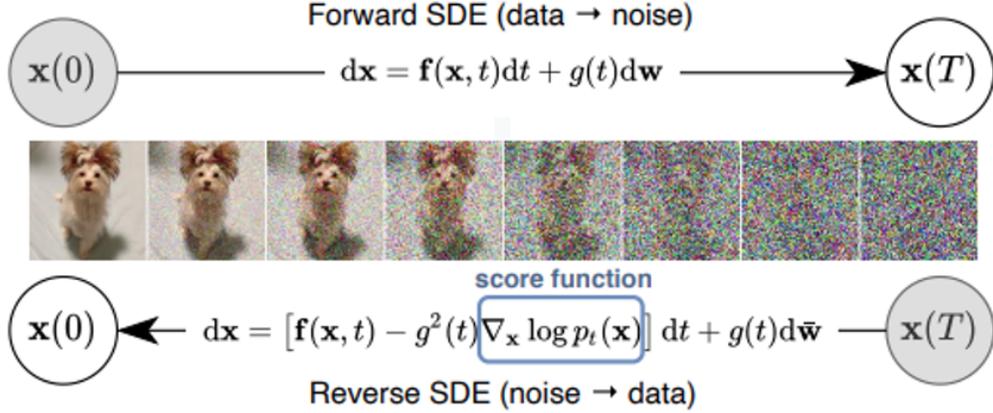


Figure 2: Diffusion models based on the reverse process. Figure from [5].

We train a Noise Conditional Score Network (NCSN), denoted by  $s_\theta(x, \sigma)$ , with a weighted sum of denoising score matching objectives:

$$\theta^* = \arg \min \sum_1^N \sigma^2 \mathbb{E}_{p_{data}(x)} \mathbb{E}_{p_{\sigma_i}(\tilde{x}|x)} [\|s_\theta(\tilde{x}, \sigma) - \nabla_{\tilde{x}} \log p_{\sigma_i}(\tilde{x}|x)\|_2^2] \quad (25)$$

Given sufficient data and model capacity, the optimal score-based model  $s_{\theta^*}(x, \sigma_i)$  matches  $\nabla_{\tilde{x}} \log p_{\sigma_i}(\tilde{x}|x)$  almost everywhere. For sampling, we run  $M$  steps Langevin MCMC to get a sample for each  $p_{\sigma_i}(x)$  sequentially:

$$x_i^m = x_i^{m-1} + \epsilon_i s_{\theta^*}(x, \sigma_i) + \sqrt{2\epsilon_i} z_i^m \quad (26)$$

where  $\epsilon_i > 0$  is the step size, and  $z_i^m$  is standard normal. This can be seen as a discretization of the reversed SDE. The above is repeated for  $i = N, N-1, \dots, 1$  in turn with  $x_N^0$  and  $x \sim \mathcal{N}(x; 0, \sigma_{\max}^2 I)$  and  $x_i^0 = x_{i+1}^M$  when  $i < N$ . As  $M \rightarrow \infty$  and  $\epsilon_i \rightarrow 0$  for all  $i$ ,  $x_1^M$  becomes an exact sample from  $p_{\sigma_{\min}} \approx p_{data}(x)$  under some regularity conditions.

## 4.2 Denoising Diffusion Probabilistic Models (DDPM) [4]

For each training data point  $x_0 \sim p_{data}(x)$ , a discrete Markov chain  $x_0, x_1, \dots, x_N$  is constructed such that

$$p(x_i|x_{i-1}) = \mathcal{N}(x_i; \sqrt{1 - \beta_i}x_{i-1}, \beta_i I),$$

and therefore  $p(x_i|x_0) = \mathcal{N}(x_i; \sqrt{\alpha_i}x_0, (1 - \alpha_i)I)$ , where  $0 < \beta_1, \dots, \beta_N < 1$ ,  $\alpha_i = \prod_{j=1}^i (1 - \beta_j)$ . In fact, this corresponds to a discretization of the following SDE (OU process)

$$dx = -\frac{1}{2}\beta(t)xdt + \sqrt{\beta(t)}dw.$$

Similar to SMLD, we can denote the perturbed data distribution as  $p_{\alpha_i}(\tilde{x}) = \int p_{data}(x)p_{\alpha_i}(\tilde{x}|x)dx$ . The noise scales are prescribed such that  $x_N$  is approximately distributed according to  $\mathcal{N}(0, I)$ . A

variational Markov chain in the reverse direction is parameterized with  $p_\theta(x_{i-1}|x_i) = \mathcal{N}(x_{i-1}; \frac{1}{\sqrt{1-\beta_i}}(x_i + \beta_i s_\theta(x_i, i)), \beta_i I)$ , and trained with a re-weighted variant of the evidence lower bound (ELBO):

$$\theta^* = \arg \min \sum_1^N (1 - \alpha_i) \mathbb{E}_{p_{data}(x)} \mathbb{E}_{p_{\alpha_i}(\tilde{x}|x)} [\|s_\theta(\tilde{x}, i) - \nabla_{\tilde{x}} \log p_{\alpha_i}(\tilde{x}|x)\|_2^2] \quad (27)$$

After solving the optimal model  $s_{\theta^*}(\tilde{x}, i)$ , samples can be generated by starting from  $x_N \sim \mathcal{N}(0, I)$  and following the estimated reverse Markov chain as below

$$x_{i-1} = \frac{1}{\sqrt{1-\beta_i}}(x_i + \beta_i s_{\theta^*}(\tilde{x}, i)) + \sqrt{\beta_i} z_i \quad (28)$$

We call this method ancestral sampling, since it amounts to performing ancestral sampling from the graphical model  $\prod_{i=1}^N p_\theta(x_{i-1}|x_i)$ . The objective is also a weighted sum of denoising score matching objectives, which implies that the optimal model,  $s_{\theta^*}(\tilde{x}, i)$ , matches the score of the perturbed data distribution,  $\nabla_{\tilde{x}} \log p_{\alpha_i}(\tilde{x}|x)$ . Notably, the weights of the i-th summand, namely  $\sigma_i^2$  and  $\alpha_i$ , are related to corresponding perturbation kernels in the same functional form:  $\sigma_i^2 \propto \mathbb{E}[\nabla_{\tilde{x}} \log p_{\sigma_i}(\tilde{x}|x)]$  and  $\alpha_i \propto \mathbb{E}[\nabla_{\tilde{x}} \log p_{\alpha_i}]$ .

### 4.3 ESTIMATING SCORES FOR THE SDE [5]

By starting from samples of  $x_T \sim p_T$  and reversing the process, we can obtain samples  $x_0 \sim p_0$ . A remarkable result from Anderson[1] states that the reverse of a diffusion process is also a diffusion process, running backwards in time and given by the reverse-time SDE:

$$d\mathbf{x} = [\mathbf{f}(\mathbf{x}, t) - g(t)^2 \nabla_{\mathbf{x}} \log p_t(\mathbf{x})] dt + g(t) d\bar{\mathbf{w}} \quad (29)$$

where  $\bar{\mathbf{w}}$  is a standard Wiener process when time flows backwards from  $T$  to 0, and  $dt$  is an infinitesimal negative timestep. Once the score of each marginal distribution,  $\nabla_x \log p_t(x)$ , is known for all  $t$ , we can derive the reverse diffusion process from Eq. 29 and simulate it to sample from  $p_0$ .

The score of a distribution can be estimated by training a score-based model on samples with score matching. To estimate  $\nabla_x \log p_t(x)$ , we can train a time-dependent score-based model  $s_\theta(x, t)$  via a continuous generalization to Eqs. 25 and 27:

$$\theta^* = \arg \min_{\theta} \mathbb{E}_t \left\{ \lambda(t) \mathbb{E}_{\mathbf{x}(0)} \mathbb{E}_{\mathbf{x}(t)|\mathbf{x}(0)} \left[ \left\| \mathbf{s}_\theta(\mathbf{x}(t), t) - \nabla_{\mathbf{x}(t)} \log p_{0t}(\mathbf{x}(t) | \mathbf{x}(0)) \right\|_2^2 \right] \right\} \quad (30)$$

Here  $\lambda : [0, T] \rightarrow R_{>0}$  is a positive weighting function,  $t$  is uniformly sampled over  $[0, T]$ ,  $x(0) \sim p_0$  and  $x_t \sim p_{0t}(x(t)|x(0))$ . With sufficient data and model capacity, score matching ensures that the optimal solution to Eq. 30, denoted by  $s_{\theta^*}(x, t)$ , equals  $\nabla_x \log p_t(x)$  for almost all  $x$  and  $t$ . As in SMLD and DDPM, we can typically choose

$$\lambda \propto 1/\mathbb{E} \left[ \left\| \nabla_{\mathbf{x}(t)} \log p_{0t}(\mathbf{x}(t) | \mathbf{x}(0)) \right\|_2^2 \right]. \quad (31)$$

## References

1. Anderson B D O. Reverse-time diffusion equation models[J]. Stochastic Processes and their Applications, 1982, 12(3): 313-326.

2. <https://www.vanillabug.com/posts/sde/>
3. Song Y, Ermon S. Generative modeling by estimating gradients of the data distribution[J]. Advances in neural information processing systems, 2019, 32.
4. Ho J, Jain A, Abbeel P. Denoising diffusion probabilistic models[J]. Advances in Neural Information Processing Systems, 2020, 33: 6840-6851.
5. Song Y, Sohl-Dickstein J, Kingma D P, et al. Score-based generative modeling through stochastic differential equations[J]. arXiv preprint arXiv:2011.13456, 2020.